



Article Technology Shocks and the Efficiency of Equity Markets in the Developed and Emerging Economies: A Global VAR Approach

Yinka S. Hammed ^{1,2} and Afees A. Salisu ^{2,3,*}

- ² Centre for Econometrics & Applied Research, Ibadan 110115, Nigeria
- ³ Department of Economics, University of Pretoria, Pretoria 0002, South Africa

Correspondence: adebare1@yahoo.com

Abstract: We tested the connection between technology shocks and the efficiency of equity markets in developed and emerging economies. We augmented the Global Vector Autoregressive (GVAR) database that covers data on 33 developed and emerging markets with the newly constructed data for technology shocks involving two variants, one with 164 countries (GTS-164), and the other, which is more region-specific. covering only Organization for Economic Co-operation and Development (OECD) countries (GTS-OECD). Our analysis was then modeled with GVAR methodology. We found that a one standard positive innovation shock to global technology (GTS-164) raises real equity prices in nearly 70% of the markets considered, and this is sustained over the forecast periods. However, the response of real equity prices to a global-specific technology shock (GTS-OECD) is rather different. While this shock resulted in the immediate rise in real equity prices, it is only transient and dissipated after the third quarter of the forecast horizon in about 85% of these markets. By implication, the efficiency of the real equity market was assured for the region-specific technology shock rather than for the more encompassing measurement that takes account of numerous markets, not minding whether these markets are developed or emerging. In sum, technological shocks seem to have greater impacts on the efficiency of developed (including Euro) markets than other markets.

Keywords: market efficiency; real equity prices; technology shock; GVAR Model

JEL Classification Codes: G14; R53; L24; B41; C32

1. Introduction

Conceptually, Efficient Market Hypothesis [hence, EMH] gives an indication that current stock prices are reflections of available information, whether private and/or public in nature (see: Saglam and Guresci 2018; Grabowski 2019). Notwithstanding this, while the EMH has been tested using various stock markets from advanced, emerging and developing countries (see for example, Yang et al. 2019; Kelikume et al. 2020), its connectivity with technological innovation remains unexplored in the literature. However, there are clear examples in the literature to suggest that a market could be efficient (see Van et al. 2013; Bulla 2015; Yang et al. 2019) or inefficient (see Nguyen and Muhammad 2011; Dong et al. 2013; Awiagah and Choi 2018; Saglam and Guresci 2018; Zaman 2019; Kelikume et al. 2020), particularly with reference to the stock market. In recent times, there is growing concern on EMH losing strength, owing especially to the technological bubble of 1995 to 2000 and the financial crisis of 2007/2008, a period which was characterized with many investors making abnormal returns (see Gilson and Kraakman 2014; McAleer et al. 2016; Kelikume et al. 2020). It is now an issue of concern on the potential roles of technology in either achieving efficient market hypothesis or the alternative. Specifically, the rate of technology advancement and its spread across various stock markets gives an indication on the nexus between the technology shock and the EMH. As a result, it is rather pertinent as to how far-reaching the impact of technological progress can further ensure EMH.



Citation: Hammed, Yinka S., and Afees A. Salisu. 2023. Technology Shocks and the Efficiency of Equity Markets in the Developed and Emerging Economies: A Global VAR Approach. *Journal of Risk and Financial Management* 16: 154. https://doi.org/10.3390/ jrfm16030154

Academic Editor: Victor Dragotă

Received: 27 January 2023 Revised: 14 February 2023 Accepted: 25 February 2023 Published: 27 February 2023



Copyright: © 2023 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/).

¹ Department of Economics, Obafemi Awolowo University, P.M.B. 13., Ile-Ife Osun 220282, Nigeria

There is evidence of studies that have discussed stock returns in relative terms via technology, though with no particular reference to international shock propagation (see Yang 2013; Kogan and Papanikolaou 2014; Garlappi and Song 2017; Prabheesh and Vidya 2018; Grabowski 2019), there are also studies on technology shock propagation in relation to exchange rate behaviour (see: Corsetti et al. 2014; Nam and Wang 2015; Miyamoto and Nguyen 2017; Klein and Linnemann 2021) and business cycle and financial friction (see Kamber et al. 2015). But, studies with specific interest on the wider implications of technology shock with respect to stock market efficiency is lacking in the literature. In other words, the discussion on the efficiency or inefficiency of markets through technology shock has not been given serious attention in the previous studies. Hence, our aim in this regard is to contribute to this discussion by investigating the role of technology shock in achieving market efficiency. For our discussion, we shall take a global look by giving consideration to a group of advanced (Euro and non-Euro countries) and emerging market economies. The idea is to present a wider implication of the shock analysis for both developed and emerging markets. If indeed the market is efficient, a positive shock to technology shocks should necessarily increase the stock volume, but not returns or price. It is believed that the investors are exposed to all available information in the markets, which would not lead to arbitrarily higher stock returns. Any other option that is rather contrary will lead to a case of weak efficiency or inefficiency.

We rely on the GVAR model for our choice of methodology. This model appropriately reflects the individual country's content of the shock propagation and also gives insight into the region-specific information. Aside, the result estimations from GVAR model analysis are usually large and include a preliminary test, impulse response functions, and main results. On the proxy for technology shock, many related studies measured it through investment specific technology shock [IST] (see, for example, Grabowski 2019) and proxy-instrumental variables from total factor productivity (see for example Klein and Linnemann 2021), but in this study, technology shock is captured through recently computed technology shock data by Sharma and Narayan (2022). The computation of this variable uses the growth of a number of patents in all of the countries considered. It equally has very large scope (more than 100 years) and the process of its computation helps to eliminate the look-ahead bias, as it is measured on a 5-year rolling window. The idea is to account for the time between the application of the patent right and the time it is often granted in many countries, where its impact within this time-frame becomes manifested in the economy.

However, these computed shock series have three components, and each has two variants: the series for the US, for the group of 164 countries (given the data availability), and the series for the group of OECD countries (12 in number). Each series appears in the original form as well as in the standardized form. As defined by the scope of this study which is essentially based on the developed and emerging market economies, we make use of the technology shock for 164 countries (hence referred to as GTS-164) and the one for OECD specific countries (hence, GTS-OECD). While computing the shocks, emphasis is much more on the detrending of growth in patents for the concerned country. In particular, the current technology shock is computed by taking the natural logarithms of the number of patents and detrend its growth rate on 5 year rolling windows. In order to compute the global technology shock, the number of patents is added for each of the 164 countries. Thereafter, its natural logarithm is taken and then detrended using a 5-year moving window. For the OECD specific technology shock, the addition of the number of patents and the detrending process is now for all 12 of the OECD countries considered in the computation (see Sharma and Narayan 2022).

We can make a summary of our findings as follows. One, we find that GTS-164 raises real equity prices in nearly 70% of the markets, and this was sustained throughout the forecast periods. This indicates that only a few markets are efficient when global technology shock is considered. In the other outcome, the response of real equity prices to GTS-OECD is rather different. While this shock results in the immediate rise in real equity prices for more than 85% of the markets, it is evidently short lived below the fourth quarter of the forecast

horizon, with an average rise of 0.72%. However, in our further analysis, the countries are grouped into various blocs where both GST-164 and GTS-OECD are considered, and the outcome reveals that many of the economic blocs could not easily absorb the GTS-164 shock except the emerging and Euro markets. As for GTS-OECD, the impact of the shock on various economic regions is transient and rarely extends beyond the second quarter of the forecast horizon. Hence, market efficiency for the equity-shock nexus is assured for region specific shock rather than the one with a more global outlook.

Apart from the foregoing, the balance of this paper entails the sections on the data and methodology, the preliminary analysis, and the result discussion. The last section presents the conclusion and the policy implications.

2. Methodology

As a way to test for the efficiency of the real equity market for the Euro, non-Euro and emerging markets, we adopt the Global Vector Auto-Regressive (GVAR) approach by Chudik and Pesaran (2013) and Smith and Galesi (2014). In this approach, we analyze the individual countries' $VARX^*(p_i, q_i)$ models across N group of developed (Euro and non-Euro) and emerging markets such that: i = 1, 2, 3, ..., N. The $VARX^*(p_i, q_i)$ model for each country and region expresses the endogenous variables as functions of both foreign and common variables. However, the GVAR model being a data base contains country-specific endogenous variables which in this case are: real GDP, inflation, real equity prices, real exchange rate, and short interest rate. In addition, the global variables are commodity prices and metal prices. We then augment these data with measures of global technology shock, of which there are of two variants: the global technology shock for the group of 164 countries (GTS-164), and the region-specific technology shock for the group of OECD countries (GTS-OECD). Given this, our GVAR framework is presented as follows:

$$x_{it} = \sum_{\lambda=1}^{p_i} \Gamma_{i\lambda} x_{i,t-\lambda} + \Pi_{i0} x_{it}^* + \sum_{\lambda=1}^{q_i} \Pi_{i\lambda} x_{i,t-\lambda}^* + \Psi_{i0} G_{it} + \sum_{\lambda}^{s_i} \Psi_{i\lambda} G_{i,t-\lambda} + \mu_{it}$$
(1)

From this equation, the variable x_{it} is a $k_i \times 1$ vector of country-specific variables denoted with i, where i runs from 1 to N (such that N = 33) in a particular period t (where t = 1, 2, 3, ..., T). Also, the variable x_{it}^* is the corresponding $k_i^* \times 1$ vector of foreign variables constructed as trade-weighted counterparts of the local variables. Hence, $x_{it}^* = \sum_{j=1}^N w_{ij} x_{jt}$ from where $\sum_{j=1}^N w_{ij} = 1$, and $w_{ii} = 0$. The external common (global) factors are represented with G_{it} and their values are repeated for all the cross-sections. Again, $\Gamma_{i\lambda}$, which runs as $\lambda = 1, 2, 3, ..., p_i$, is a $k_i \times k_i$ matrix of unknown parameters for domestic variables; $\Pi_{i\lambda}$ (where $\lambda = 0, 1, 2, ..., q_i$) is a $k_i^* \times k_i^*$ matrix of unknown parameters for the lag of foreign variables, and $\Psi_{i\lambda}$ (where $\lambda = 0, 1, 2, ..., s_i$) is an $r_i \times r_i$ matrix of unknown parameters for the lag of global variables which are repeated for all the cross-sections; while μ_{it} is a vector (i.e., $k_i \times 1$) of error terms. Both Π_{i0} and Ψ_{i0} are the parameters for the current periods as defined for variables above. However, we factor all the foreign and common factor variables as weakly exogenous in this context. As shown in Equation (1), we stack the estimated country-specific model to form a large GVAR model from which the effect of the global technology shock on international real equity prices is formed.

3. Data and Preliminary Analyses

There are two datasets combined in this study. The first is the traditional GVAR dataset that contains both the domestic and foreign variables. On this, we refer our readers to the GVAR toolbox of Smith and Galesi (2014) for the original data as well as the recently updated version by Mohaddes and Raissi (2020)², which covers the period up to the end of 2019. The second dataset which covers the technology shock data is incorporated into the GVAR model as a global variable. The construction of technology shock that this study employs follows some specific steps which principally involve the detrending of growth in patents for the concerned country (see Sharma and Narayan (2022) for technical details).³

In particular, the current technology shock is estimated by taking the natural logarithms of the number of patents and detrending its growth rate on 5-year rolling windows. The 5-year window accounts for the periods that often lie between the application for the patent right and the period when it is granted in various countries of consideration⁴. This process of detrending as applied in deriving the shocks essentially helps to eliminate the problem associated with look-ahead bias. A similar approach is followed to derive the global technology shock, where the number of patents is added for each of the 164 countries for which data are available. Thereafter, its natural logarithm is taken and then detrended using a 5-year moving window. For the OECD-specific technology shock, the addition of the number of patents and the detrending process is now for only 12 OECD countries (see Sharma and Narayan 2022 and other papers as cited in the paper in lieu of construction of the shock). Although the technology shock series has a broader scope (more than 100 years for the least countries), our scope is limited to the periods between 1980 and 2018. The start date is defined by the data on real stock prices (as well as other variables in the GVAR model), and the end date is defined by the series on technology shock. However, while the real equity prices are quarterly data, that of the technology shock is annual. It is believed that market efficiency will hold when stock prices have no significant response to the innovation arising from technology shock and the reverse will be assured when the opposite takes place.

The background information about the variables is provided in Table 1. This information is specifically for the risk adjusted mean and the coefficient of variations for the real equity prices across the group of developed and emerging market economies and that of the global technology shock and specific OECD shock. While the global shock is derived for the group of 164 countries, that of OECD shock is mainly for 12 OECD countries. The risk adjusted mean for the real equity prices across the markets ranges between 0.487 and 7.367. In all, only two cases have values below 1 (Argentina and Finland). By implication, the average values for the real equity are higher than the level of dispersion in virtually all the markets, suggesting the possibility of low risk in the markets. As for the coefficient of variation, the values are found to be below 100% except for very few cases. In fact, in about 62% of the markets, the standard deviation is only about 40% of the mean, giving an indication of relatively lower dispersion. For the information about various economic blocs, both the risk adjusted mean and coefficient of variation are normal, though the real equity for emerging markets has a relatively higher level of dispersion (the standard deviation being 42.2% of the mean). However, the global technology shock and OECD specific shock have higher levels of volatilities (the standard deviation is almost twice the mean of the two shocks series). By implication, there is higher variability among the computed series. We further align the real equity prices in all the countries and the bloc regions with technology shock series in a graphical representation (see Appendix A and Figures A1 and A2). From these graphs, we infer an inverse relationship between the equity prices and the shock component series. This gives an implication that positive technology shocks command higher prices in many of the markets.

Table 1. Descriptive Statistics.

Real Equity Prices	Mean	Std. Dev.	RAM	CoV
Argentina	0.266	0.546	0.487	2.053
Australia	2.091	0.459	4.556	0.219
Austria	1.442	0.647	2.229	0.449
Belgium	1.543	0.560	2.758	0.363
Canada	1.848	0.584	3.164	0.316
Chile	2.094	1.267	1.653	0.605
Finland	0.741	0.957	0.774	1.292

Real Equity Prices	Mean	Std. Dev.	RAM	CoV
France	2.314	0.649	3.567	0.280
Germany	1.462	0.523	2.797	0.357
India	1.559	0.821	1.900	0.526
Italy	2.613	0.421	6.207	0.161
Japan	1.316	0.363	3.626	0.276
South Korea	0.629	0.599	1.049	0.953
Malaysia	1.573	0.436	3.608	0.277
Netherlands	1.772	0.640	2.771	0.361
Norway	1.839	0.733	2.509	0.399
New Zealand	0.451	0.254	1.774	0.564
The Philippines	1.481	1.130	1.310	0.763
South Africa	3.899	0.656	5.943	0.168
Singapore	0.925	0.323	2.861	0.350
Spain	0.682	0.674	1.012	0.988
Sweden	2.380	0.894	2.662	0.376
Switzerland	1.445	0.755	1.914	0.522
Thailand	1.461	0.717	2.036	0.491
United Kingdom (UK)	3.169	0.430	7.367	0.136
United States (US)	1.817	0.662	2.746	0.364
Euro	1.571	0.592	2.653	0.377
Non Euro	1.680	0.563	2.987	0.335
Developed Less Euro	1.568	0.503	3.116	0.321
Emerging Markets	1.620	0.683	2.372	0.422
	Technolo	gy Shock		
GTS-164	0.081	0.135	0.599	1.669

Table 1. Cont.

Note: RAM is the risk adjustment mean, and it is computed as the mean divided by the corresponding standard deviation; CoV is the coefficient of variation, and it is computed as the standard deviation divided by the corresponding mean. However, all the variables are expressed in natural logs. Also, GTS-164 is Global Technology Shock derived for the group of 164 countries, while GTS-OECD is the shock computed for the OECD countries.

0.151

0.588

1.701

0.089

4. Results and Discussion

GTS-OECD

Our results are divided into two. In the first part, we use GTS-164 (i.e., the global technology shock which captures the aggregate shock for 164 countries), and in the second, we employ GTS-OECD (i.e., the global technology shock for the group of 12 OECD countries with the inclusion of the US). Figures 1 and 2 give the response of real equity to a positive standard deviation of GTS-164 for both country-level analysis and the region-specific evaluation. In Figure 1, where we present the country-level analysis, the shock to the positive standard deviation of the GTS-164 results in overall fall in the real equity prices in about 70% of the markets considered, and we further find this response to be instantaneous in about 46% of the markets and delayed in 23% of the total countries considered. However, while the shock impact is sustained for the significant number of the markets with an average rise in real equity of 0.62%, it is rather not persistent for the few others, and the real equity price, in this case, only rises by 0.24% on average (see Figure 1).

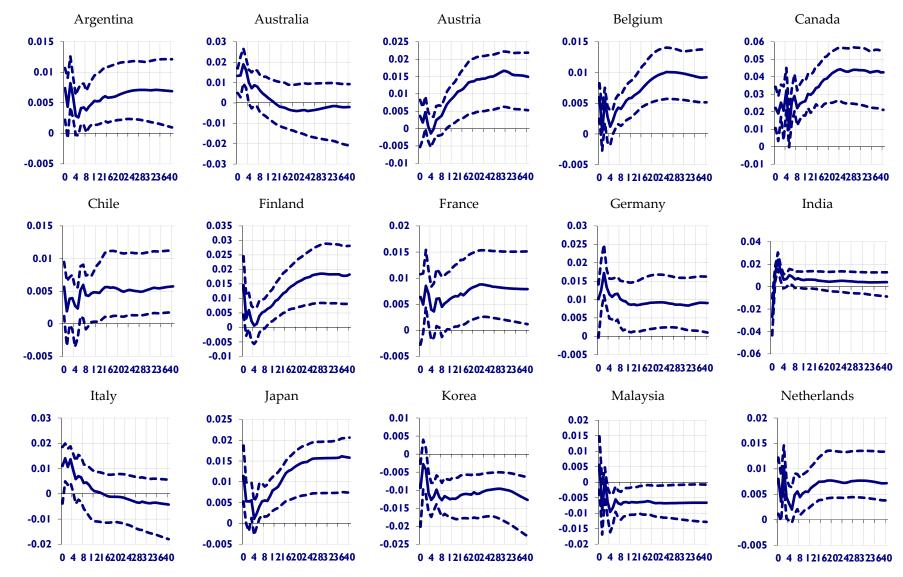


Figure 1. Cont.

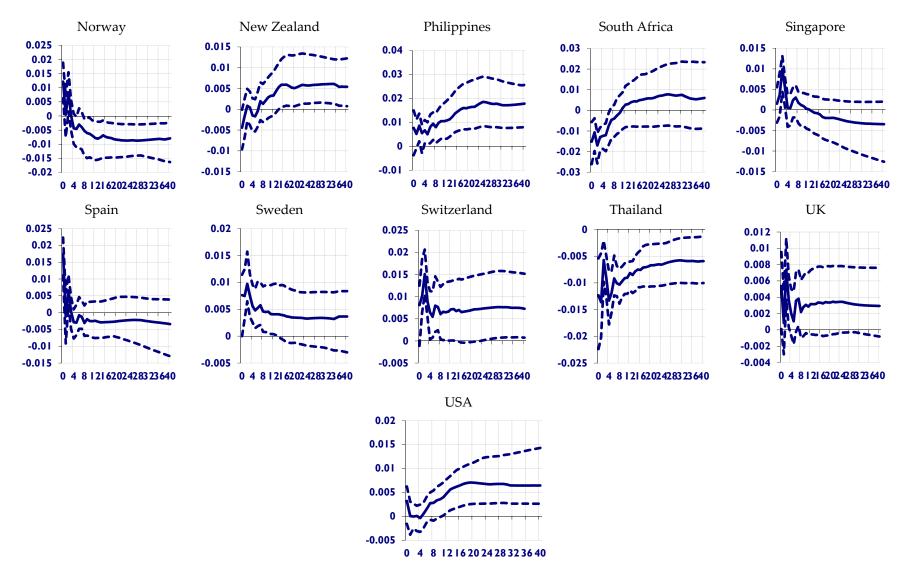


Figure 1. Country-level Impulse Response Functions of Real Equity Prices to a One Standard Deviation Positive Global Technology Shock (GTS – 164).

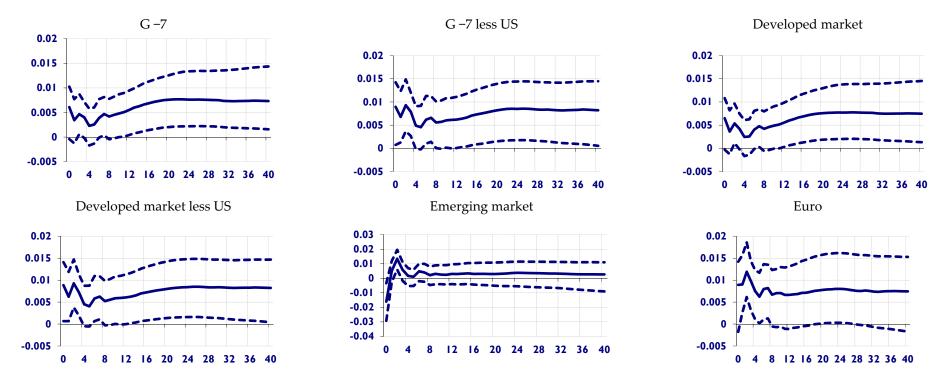


Figure 2. Group Impulse Response Functions of Real Equity Prices to a One Standard Deviation Positive Global Technology Shock (GTS - 164).

As emphasized before, for efficiency to hold for any market, stock prices should necessarily not be influenced by any factors other than the available information in the market. Meanwhile, a further assessment of the propagation of this shock gives some indication that many Euro countries are rather inefficient at absorbing the global technology shocks (the impact is sustained with an average rise of 0.74%), while other developed and emerging markets behave relatively better. In sum, the shock impact only renders very few markets as being efficient, while greater numbers of others are not. Surprisingly, the US economy cannot absorb this shock, with the initial response being very sluggish for up to 10 quarters of the forecast horizon, but sustained later. This could possibly suggest the ability of the US economy to withstand any external shock at any period when it occurs. In Figure 2, the countries are categorized into various economic blocs for further analysis. The results give further support to the analysis at the country level. The developed market responds to the Global technology shock with an immediate fall in real equity prices when the US is not considered. However, the impact becomes delay with the inclusion of US, where real equity prices rise by an average of 0.7% and at the same time having the impact being sustained throughout the forecast period. The influence of the US in this nexus becomes that much more established. The response of real equity for the Euro and emerging markets to the shock innovation is evident with the immediate rise in real equity. While this rise (on average) is 0.29% for the emerging markets, it is higher for the European bloc (approximately 0.77%), and it was further found to be transient for the two markets. Thus, by categorizing the countries into blocs, the Euro market and the emerging markets are efficient, as the shock impact could not be sustained. In all, the efficient market hypothesis does not hold when global technology shock is considered. The arising implication is that stock prices are capable of being influenced when they are exposed to global shock variables. Hence, the impact of global technology shock in changing the stock prices is rather obvious. While the efficiency market hypothesis does not hold in this regard, it suggests that investors' rationality or irrationality could rather give way to adaptive market efficiencies for the global technology shock. This is more likely as changes in stock prices are found to be associated with significant changes in global technology shocks.

As a way to make this finding more engaged, we give consideration to global technology shock for a group of OECD countries. This is done to establish whether this shock impact will exhibit different trends when the global technology shock becomes more regionspecific. The response of various real equity markets to the OECD global technology shock, however, gives some salient information on how real equity prices of different countries could be prone to the technology shock that is rather region-specific. Figures 3 and 4 present this information for all the developed and emerging markets that are considered in this study. In Figure 3, we observed an immediate rise in real equity prices for nearly 85% of the economies under investigation. However, this response is transient and could rarely last for up to three quarters of the forecast horizon. By implication, while a more global measure of technology shock renders many individual markets inefficient, the shock impact of a region-specific technology is rather different. For the markets where the shock impact is transient, the real equity prices rise by an average of 0.72%, and higher where it is persistent. We now find that many Euro countries easily absorb this shock, and the same scenario is evident for other developed and emerging markets. In addition, the US economy further established its strong stance by being resilient to this shock impact. Therefore, the efficiency of the markets for these economies given the consideration of OECD global technology shock is assured for nearly all the countries under consideration.

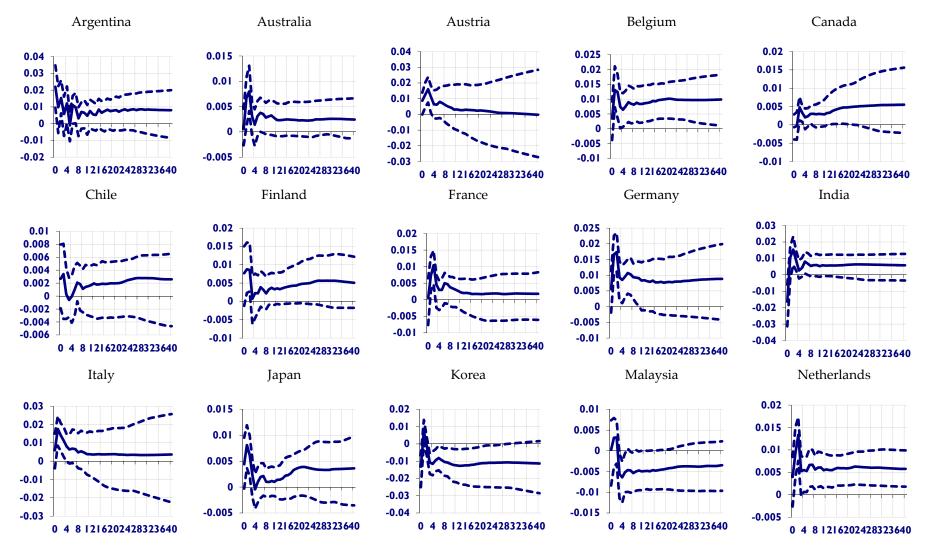


Figure 3. Cont.

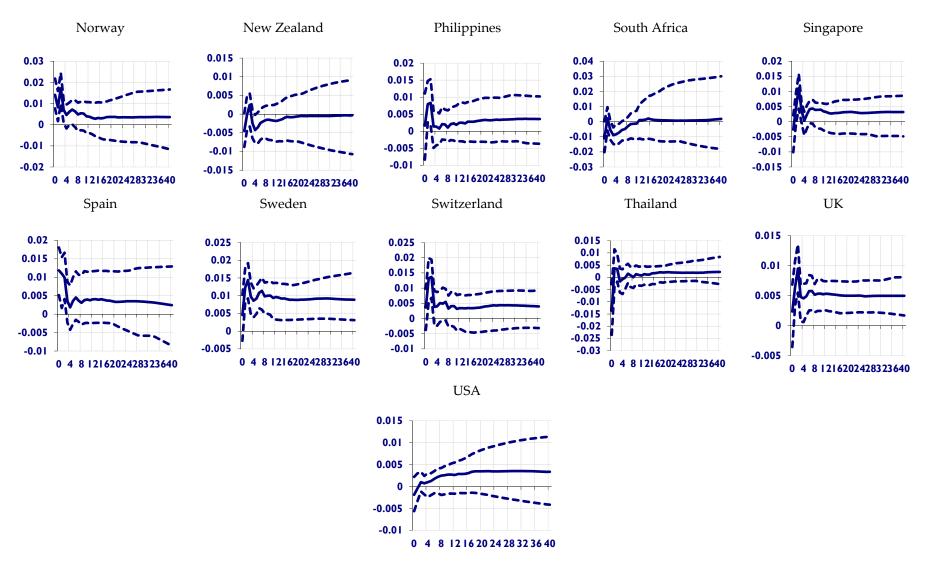


Figure 3. Group Impulse Response Functions of Real Equity Prices to a One Standard Deviation Positive US Global Technology Shock (GTS – OECD).

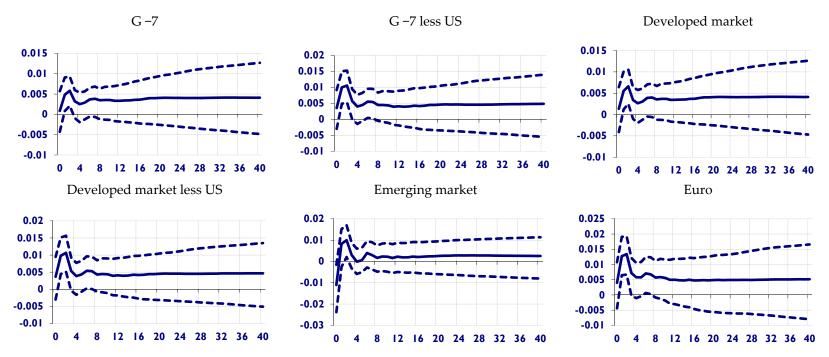


Figure 4. Group Impulse Response Functions of Real Equity Prices to a One Standard Deviation Positive Global Technology Shock (GTS – OECD).

Accounting for the global technology shock from the perspective of OECD economies provides more insight on other developed and emerging markets' efficiencies. Thus, its negligence could have led to the wrong conclusion. We then argue that where the global technology shock is limited to few countries that share economic values with developed and emerging markets, the impact on real equity prices is likely to be transient and the efficiency of the markets will be established. However, measuring the impact of a global technology shock by aggregating the data of many countries without referencing them as either developed or emerging markets could be misleading. In Figure 4, where the countries are grouped into different economic blocs, the response of real equity prices to the positive innovation of OECD technology shock is only transient, and rises by 0.53%(the most in the second forecast horizon) for the G-7 countries. However, when the US is excluded, the impact further spread to the sixth and seventh periods of the forecast horizon, with an average increase in real equity price of 0.71%. As a result, the presence of US in absorbing this shock is further established. Furthermore, the response of real equity prices of emerging markets to the shock is short-lived (it only occurs in the second quarter, with an average increase of 0.9%), while that of Euro markets is only transient and dissipates after the seventh quarter of the forecast horizon (an average increase of 0.98%). Thus, while the global technology shock is only absorbed by Euro and emerging markets, the OECD shock is easily absorbed by all of the blocs. Hence, market efficiency for the equity-shock nexus is assured for a region-specific shock rather than the one with a more global outlook. In other words, the efficient market hypothesis is found to hold for all the markets considered when the technological shock is region-specific. By implication, this type of technology shock has little or no impact on the stock prices of the countries under consideration. In this regard, the efficient market hypothesis explains the transient and insignificant influence of technology shock on stock prices. Thus, while the adaptive market hypothesis holds for global technology shock, EMH is confirmed for region-specific shocks.

5. Conclusions

We employed recently constructed technology shock data by Sharma and Narayan (2022) together with real equity prices for a group of emerging and developing countries in a GVAR model. The data were of three variants: the global technology shock that covers 164 countries (GTS-164), the region specific one for 12 OECD countries (GTS-OECD), and the specific one for the US. We are, however, more focused on testing the efficiency of real equity markets for Euro, non-Euro and emerging markets with data on global technology shock (GTS –164) and the specific regional shock (GTS –OECD) that spans from 1980Q1 to 2018Q4. Our choice of GVAR model lies in its efficacy to reveal global specific information of the variables under consideration without losing sight of the individual countries' content.

Our findings can be summarized as follows. One, GTS-164 raises real equity prices in nearly 70% of the markets, and this was sustained throughout the forecast periods. This indicates that only a few markets are efficient when global technology shock is considered. In the other outcome, the response of real equity prices to GTS-OECD is rather different. While this shock results in the immediate rise in real equity prices for more than 85% of the markets, it is evidently short lived below the fourth quarter of the forecast horizon, with an average rise of 0.72%. However, in our additional analysis, the countries are grouped into various blocs where both GST-164 and GTS-OECD are considered, and the outcome reveals that many of the economic blocs could not easily absorb the GTS-164 shock except the emerging and Euro markets. As for GTS-OECD, the shock impact on various economic regions is transient and rarely extends beyond the second quarter of the forecast horizon.

The implication arising is that the GTS-164 is more intense in its impact on real equity prices across many markets, which suggests that the real equity markets in those countries (including US) are inefficient. However, the outcome from GTS-OECD shocks indicates that many of the markets and various economic blocs (including the Euro region) as considered in this study are efficient. It thus suggests that the efficiency of real equity market is associated with region-specific technology shock but rather not for the alternative

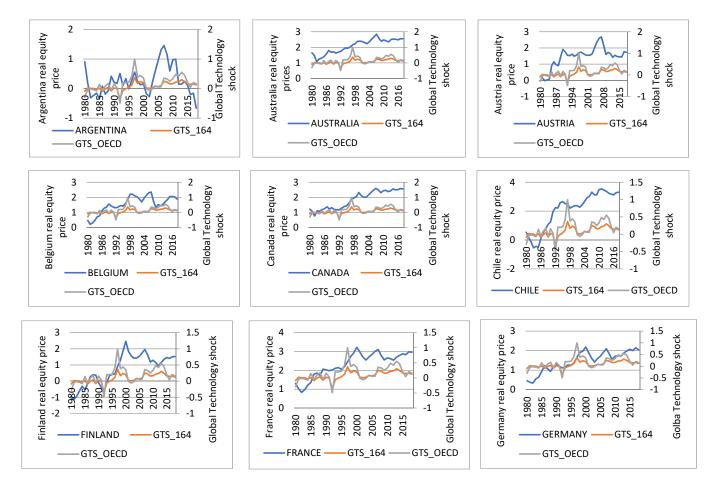
measures that account for numerous markets with no consideration on their relationship with developed and emerging countries. In such a situation when technology shock is particularly global, adaptive market hypotheses play an important role in explaining the changes in stock prices rather than the Efficient Market Hypothesis.

Author Contributions: Conceptualization, A.A.S. and Y.S.H.; Methodology, A.A.S. and Y.S.H.; Software, A.A.S. and Y.S.H.; Validation, A.A.S.; Formal Analysis, A.A.S. and Y.S.H.; Resources, A.A.S. and Y.S.H.; Data Curation, Y.S.H.; Writing—original draft preparation, Y.S.H.; Writing—reviewing and editing, A.A.S.; Supervision, A.A.S. All authors have read and agreed to the published version of the manuscript.

Funding: This research received no external funding.

Data Availability Statement: The GVAR data used in this article can be found on the following link: http://www.econ.cam.ac.uk/people-files/emeritus/mhp1/GVAR/GVAR.html, accessed on 26 January 2023. However, data on technology shocks can be made available upon request from the corresponding author.

Conflicts of Interest: The authors declare no conflict of interest.



Appendix A

Figure A1. Cont.

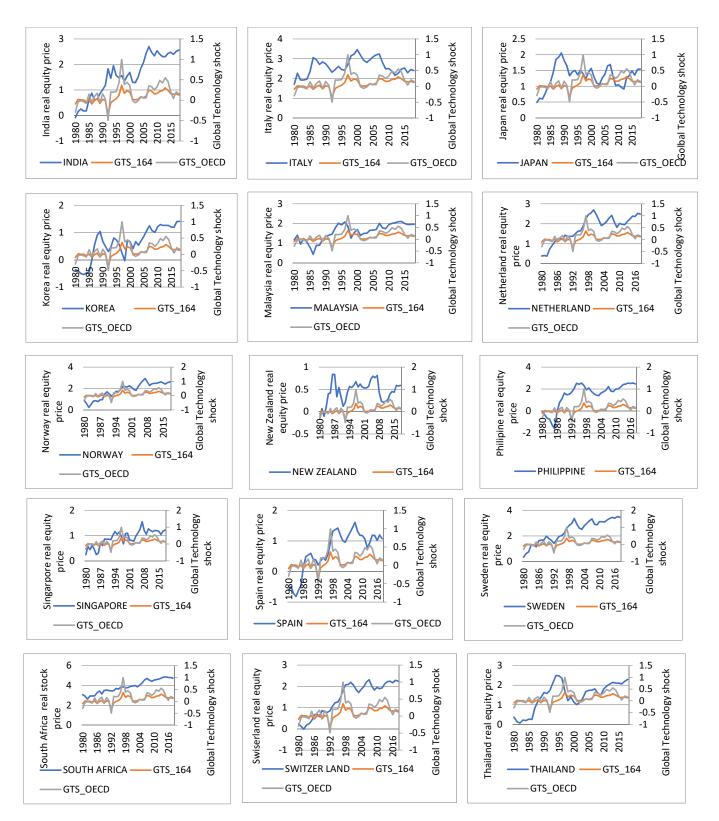


Figure A1. Cont.

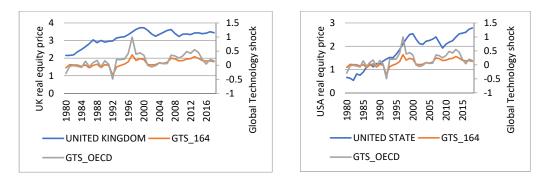


Figure A1. Country-level co-movement of real equity prices and technology shock.

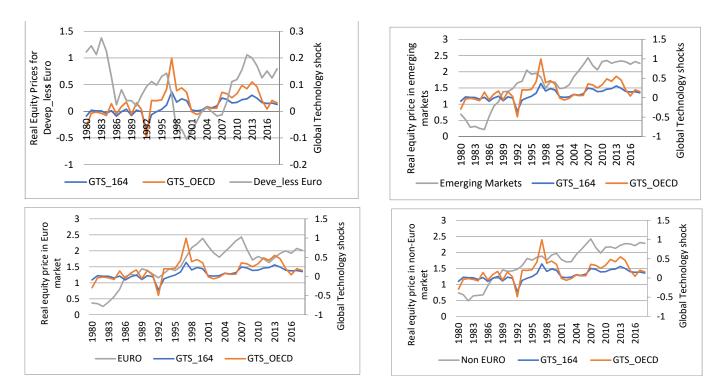


Figure A2. Region level co-movement of real equity price and technology shock. Note: GTS–164 is the global technology shock for the group of 164 countries while GTS—OECD is for the shock computed for the group of OECD countries. Also, EURO indicates the European countries, Non–EURO indicates the Non–European countries, and Deve_less Euro indicates the other developed countries that are non–Euro.

Notes

- ¹ w_{ij} is the weighting matrix obtained from the IMF Direction of Trade flows data.
- ² See link to the data at: http://www.econ.cam.ac.uk/people-files/emeritus/mhp1/GVAR/GVAR.html (accessed on 3 January 2023).
- ³ We spliced the TS-related data into quarterly observations in order to maintain the same frequency as the GVAR datasets. Since we do not expect the TS to change frequently, maintaining a constant average over four quarters in a year aligns with the expected behaviour of technological development.
- ⁴ For a detailed analysis and other associated studies, see Hsu (2009) and Hirshleifer et al. (2013).

References

- Awiagah, Raphael, and Sung Sup Brian Choi. 2018. Predictable or Random?—A Test of the weak-form efficient market hypothesis on the Ghana stock exchange. *Journal of Finance and Economics* 6: 213–22. [CrossRef]
- Bulla, Dennis. 2015. Random walk hypothesis in emerging stock markets: Evidence from the Nairobi securities exchange. International Journal of Economics, Finance and Management 4: 99–104.

- Chudik, Alexander, and Mohammad Hashem Pesaran. 2013. Econometric analysis of high dimensional VARs featuring a dominant unit. *Econometric Reviews* 32: 592–649. [CrossRef]
- Corsetti, Giacarlo, Luca Dedola, and Sylvain Leduc. 2014. The international dimension of productivity and demand shocks in the US economy. *Journal of the European Economic Association* 12: 153–76. [CrossRef]
- Dong, Huijian, Bowers Mary Helen, and William. R. Latham. 2013. Evidence on the efficient market hypothesis from 44 global financial markets. *Economics Research International* 2013: 2–11. [CrossRef]
- Garlappi, Lorenzo, and Zhongzhi Song. 2017. Can investment shocks explain the cross section of equity returns? *Management Science* 63: 3829–48. [CrossRef]
- Gilson, Ronald J., and Reinier Kraakman. 2014. Market efficiency after the financial crisis: It's still a matter of information costs. *Virginia Law Review* 100: 313–75. [CrossRef]
- Grabowski, David. 2019. Technology, Adaptation and the Efficient Market Hypothesis. *Adaptation and the Efficient Market Hypothesis*. Available online: https://papers.ssrn.com/sol3/papers.cfm?abstract_id=3649446 (accessed on 5 January 2023). [CrossRef]
- Hirshleifer, David, Po-Hsuan Hsu, and Dongmei Li. 2013. Innovative efficiency and stock returns. *Journal of Financial Economics* 107: 632–54. [CrossRef]
- Hsu, Po-Hsuan. 2009. Technological innovations and aggregate risk premiums. Journal of Financial Economics 94: 264–79. [CrossRef]
- Kamber, Günes, Christie Smith, and Christoph Thoenissen. 2015. Financial frictions and the role of investment-specific technology shocks in the business cycle. *Economic Modelling* 51: 571–82. [CrossRef]
- Kelikume, Ikechukwu, Evans Olaniyi, and Faith A. Iyohab. 2020. Efficient market hypothesis in the presence of market imperfections: Evidence from selected stock markets in Africa. *International Journal of Management, Economics and Social Sciences (IJMESS)* 9: 37–57. [CrossRef]
- Klein, Mathias, and Ludger Linnemann. 2021. Real exchange rate and international spillover effects of US technology shocks. *Journal of International Economics* 129: 103414. [CrossRef]
- Kogan, Leonid, and Dimitris Papanikolaou. 2014. Growth opportunities, technology shocks, and asset prices. *The journal of finance* 69: 675–718. [CrossRef]
- McAleer, Michael, John Suen, and Wing-Keung Wong. 2016. Profiteering from the dot-com bubble, subprime crisis and Asian financial crisis. *The Japanese Economic Review* 67: 257–79. [CrossRef]
- Miyamoto, Wataru, and Thuy Lan Nguyen. 2017. Understanding the cross-country effects of U.S. technology shocks. *Journal of International Economics* 106: 143–64. [CrossRef]
- Mohaddes, Kamiar, and Mehdi Raissi. 2020. *Compilation, Revision and Updating of the Global VAR (GVAR) Database, 1979Q2–2019Q4.* Cambridge: Judge Business School, University of Cambridge.
- Nam, Deokwoo, and Jian Wang. 2015. The effects of surprise and anticipated technology changes on international relative prices and trade. *Journal of International Economics* 97: 162–77. [CrossRef]
- Nguyen, Chu Vo, and Mahboob Ali Muhammad. 2011. Testing the weak efficient market hypothesis using Bangladeshi panel data. Banks & Bank Systems 4: 11–15.
- Prabheesh, Kumar Pahdan, and C. T. Vidya. 2018. Do business cycles, investment-specific technology shocks matter for stock returns? *Economic Modelling* 70: 511–24. [CrossRef]
- Saglam, Yağ;mur, and Gülçin Guresci. 2018. A Random Walk of Stock Prices in Visegrad Group: Efficient Market Hypothesis. In *Efficiency in Business and Economics. Springer Proceedings in Business and Economics.* Cham: Springer, pp. 175–85. [CrossRef]
- Sharma, Susan Sunila, and Paresh Kumar Narayan. 2022. Technology shocks and stock returns: A long-term perspective. *Journal of Empirical Finance* 68: 67–83. [CrossRef]
- Smith, L. Vanessa, and Alesandro Galesi. 2014. GVAR Toolbox 2.0. Available online: https://sites.google.com/site/gvarmodelling/gvar-toolbox (accessed on 5 January 2023).
- Van, Dorathea Heerden, Jose Rodrigues, Dale Hockly, Bongani Lambert, Taljard Tjaart, and Andrew Phiri. 2013. Efficient Market Hypothesis in South Africa: Evidence from a Threshold Autoregressive (TAR) Model. Munich Personal RePEc Achieve (MPRA), MPRA Paper No. 50544. Available online: https://mpra.ub.uni-muenchen.de/50544/ (accessed on 5 January 2023).
- Yang, Fan. 2013. Investment shocks and the commodity basis spread. Journal of Financial Economics 110: 164-84. [CrossRef]
- Yang, Yan-Hong, Ying-Hui Shao, Hao Lin Shao, and H. Eugene Stanley. 2019. Revisiting the weak-form efficiency of the EUR/CHF exchange rate market: Evidence from episodes of different Swiss franc regimes. *Physica A: Statistical Mechanics and Its Applications* 523: 734–46. [CrossRef]
- Zaman, Shafir. 2019. Weak form market efficiency test of Bangladesh Stock Exchange: An empirical evidence from Dhaka Stock Exchange and Chittagong Stock Exchange. *Journal of Economics, Business and Accountancy Ventura* 21: 285–91. [CrossRef]

Disclaimer/Publisher's Note: The statements, opinions and data contained in all publications are solely those of the individual author(s) and contributor(s) and not of MDPI and/or the editor(s). MDPI and/or the editor(s) disclaim responsibility for any injury to people or property resulting from any ideas, methods, instructions or products referred to in the content.