



1 Supplementary materials

2 1. The position of MEG channels

- 3 To determine which brain areas are related to the MEG channels selected by CBPT in our results,
- 4 we referred to Figure S1 [1, 2], which shows Elekta Neuromag MEG channel groupings in different
- 5 brain areas.



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Figure S1. Presentation of the MEG sensor grouping in five brain areas

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9 2. Assessing the stimuli

After the MEG measurement, subjects were requested to rate each picture stimuli using the Self-Assessment Manikin (SAM; [3]). SAM provides a seven-point scale indicating arousal (1 to 7, relaxed to excited) and valence (1 to 7, pleasant to unpleasant) levels. Comparing the mean rating of subjects for each picture category using the t-test we found significant differences which are shown in Figure S2 and Table S1.

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 Table S1. Results of comparing mean ratings of subjects for each picture category in this study.

	Comparison	p-value
Va	Pl vs. Ne	10 ⁻²⁹
alen	Un vs. Ne	10 ⁻³⁸
ce	Un vs. Pl	10^{-50}
A	Pl vs. Ne	10^{-26}
rous	Un vs. Ne	10 ⁻³⁷
al	Un vs. Pl	10 ⁻¹⁹

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Pl: Pleasant, Ne: Neutral, Un: Unpleasant.



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Figure S2. Mean rating of arousal and valence over subjects for each picture category. Circles show the mean ratings of arousal and valence across all subjects for each picture category (pleasant (Pl): blue circles; neutral (Ne): green circles; unpleasant (Un): red circles) selected from the International Affective Picture System (IAPS). The horizontal axis shows valence (pleasant-to-unpleasant) values and the vertical axis shows arousal (low-to-high) values from a 1-7 scale. Unpleasant pictures were rated as the most arousing pictures compared to pleasant and neutral stimuli. Pleasant stimuli were rated as the lowest valence pictures and unpleasant stimuli as the highest valence pictures.

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27 3. Classification method

28 We performed a classification algorithm on features based on the selected channels and 29 time-intervals by the CBPT. We took the mean values of the power spectrum over all trials of each 30 emotion category and over frequencies of a particular frequency-band and over the selected 31 time-windows by CBPT as observations for each subject. Thus in each frequency-band, each subject 32 provided three vector-valued observations for each frequency-band (5 bands): one for pleasant, one 33 for neutral, and one for unpleasant. Each observation in each comparison incorporates a specific 34 number of elements which were the magnetometers selected by CBPT. Therefore, for both two 35 comparisons (pleasant and unpleasant vs. neutral) we compiled special feature sets based on 36 selected channels and selected time-windows from CBPT with a dimensionally of 34 samples (17 37 subjects × 2 conditions) × number of selected channels by CBPT in each time-windows.

After feature sets definition, we applied the regularized logistic regression with the most popular penalty, the Least Absolute Shrinkage and Selection Operator (LASSO; [4]), to the feature-sets. LASSO is a highly common classification method and because it selects the most reliable feature-subsets by forcing some regression coefficients to zero, it provides feature-subset selection together with high classification accuracies [5]. Here, we defined the response variable of 43 the logistic regression by one for emotional conditions (pleasant and unpleasant), and zero for the 44 neutral condition. Accordingly, the probability of being in an emotional state (class 1) for the k^{th} 45 subject is estimated by Equation (1):

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$$\pi_k = p(y_k = 1 | x_k) = \frac{\exp(\beta_0 + \sum_{m=1}^M x_{km}^T \beta_m)}{1 + \exp(\beta_0 + \sum_{m=1}^M x_{km}^T \beta_m)} \qquad k = 1, 2, \dots, K$$
(1)

47 where $y_k \in \{0, 1\}$ is a vector with *k* elements, x_k are associate vectors with *M* predictors, β_m are 48 the regression coefficients, and β_0 is the intercept. The ratio of the probability of class 1 divided by 49 the probability of the class 0 (which is called odds function) [6] and then logit transformation is 50 estimated by Equation (2):

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$$Odds = \frac{\pi_k}{1 - \pi_k} = \exp \left(\beta_0 + \sum_{m=1}^M \mathbf{x}_{km}^T \beta_m\right) \quad \Rightarrow Ln\left(\frac{\pi_k}{1 - \pi_k}\right) = \beta_0 + \sum_{m=1}^M \mathbf{x}_{km}^T \beta_m \tag{2}$$

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53 To estimate the β_m and β_0 we have to calculate log-likelihood function, which is defined as 54 Equation (3):

$$l(\beta_0,\beta) = \sum_{k=1}^{K} \{y_k \ln(\pi_k) + (1-y_k) \ln(1-\pi_k)\}$$
(3)

55 which estimates the regression coefficients β_m based on both test and train data and leads to lower 56 variances and better classification accuracies. The tuning parameter, λ , controls the shrinkage. This 57 means When $\lambda = 0$, the penalty term has no effect, and as $\lambda \to \infty$, the impact of the penalty term 58 grows. However, to minimizing the whole equation, the regression coefficients will approach zero in 59 the case of $\lambda \to \infty$. This forces some of the coefficients to become zero. Thus, selecting the best value 60 for λ is critical, and this is usually done using cross-validation, which involves testing a number of 61 different values and selecting one that minimizes cross-validated error on test data [6, 7]. In this 62 study, we performed leave-one-subject-out cross-validation to selected optimal λ for each feature 63 sets. Then the classification performances using this λ for each feature sets were evaluated by 64 accuracy (the ratio of correctly classified samples to their total number). The accuracies were also 65 assessed performing 1000 17-fold-stratified cross-validations and 95%- confidence intervals of them 66 were reported. It should be noted that we used 17 folds because we had 17 subjects.

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